Supplemental Appendix for

Bespoke Realized Volatility: Tailored Measures of Risk for Volatility Prediction

by Andrew J. Patton and Haozhe Zhang

This version: December 30, 2022

This appendix contains six subsections, presenting additional details and analyses relevant to the main paper.

S.1. Data cleaning details

We use trade data from the NYSE Trade and Quote (TAQ) database. We follow Barndorff-Nielsen, Hansen, Lunde and Shephard (2009) to clean this data, using the following rules:

- 1. Keep only entries with time step between 9:30 am to 4:00 pm (when the exchange is open).
- 2. Delete entries with zero transaction prices.
- 3. Retain entries originating from NYSE and NASDAQ only.
- 4. Delete entries with corrected trades (CORR $\neq 0$).
- 5. Delete entries with unusual sale condition (COND with letter code except for letters E and F).
- 6. We use the median price if multiple entries have the same time stamp.

S.2. Simulation study

In this section we simulate data to illustrate the ability of the proposed Cubic HAR model to recover the true coefficients. We assume that high-frequency returns, $r_{i,t}$, are i.i.d. N(0,1/M), where M is the number of high-frequency returns each trade day. We combine these high frequency returns to obtain realized variances using time-of-day (TOD) weights, ω_i , as in equation 7 of the main paper, and we use a HAR model to simulate the time series dynamics of RV.

$$RV_t = \beta_0 + \beta_d \sum_{i=1}^{M} \omega_i r_{i,t-1}^2 + \beta_w \frac{1}{4} \sum_{i=2}^{5} \sum_{i=1}^{M} \omega_i r_{i,t-j}^2 + \beta_m \frac{1}{16} \sum_{i=6}^{21} \sum_{i=1}^{M} \omega_i r_{i,t-j}^2 + \epsilon_t$$
 (S.1)

where $\epsilon_t \sim N(0,1)$. We set $\beta_0 = 0$, $\beta_d = 0.4$, $\beta_w = 0.3$, $\beta_m = 0.3$, and M = 78. We impose that ω_i follows the empirical TOD weights presented in Figure 6. Then we use the cubic HAR model with the exact same set-up (hyperparameter grid and estimation algorithm) to estimate the coefficients on the simulated data. Figure S.1 shows that when the true data generating process indeed is a linear combination with the TOD weights, our cubic HAR model will recover the true weights almost perfectly. (As the true weights in this simulation are not a smooth function of time, the cubic spline model clearly cannot obtain a perfect fit.) This simulation results give us confidence that the proposed Cubic HAR has enough estimation power to recover the optimal weighting scheme for realized variance forecasting.

Cubic Spline HAR Raw Weights: Daily Lag Cubic Spline HAR Raw Weights: Weekly Lag Estimated Weights
True Weights
Equal Weights Estimated Weights
 True Weights --- Equal Weights 0.40 0.5 Raw Daily Lag Weights Raw Weekly Lag M 0200 0.15 0.1 10:10 10:30 10:50 11:10 11:30 11:50 12:30 12:50 12:50 13:10 13:50 11:50 12:10 12:30 12:50 13:10 13:50 13:50 14:10 Cubic Spline HAR Raw Weights: Monthly Lag Cubic Spline HAR Raw Weights on Simulated Data Estimated Weights
True Weights
Equal Weights Daily Lag Monthly Lag Weekly Lag 0.5 0.25 Raw Monthly Lag Weights Raw Weights 00 0.2 0.10 0.05

Figure S.1: Cubic HAR vs True Weights: Simulated Data

Note: This figure compares the estimated optimal weights (solid line) based on simulated data based with the true weights (dash-dotted line).

S.3. Alternative target for shrinkage

In the analysis in Section 3.2 of the main paper, we consider shrinkage methods (Ridge, LASSO and elastic net) that shrink the estimated parameters towards zero. This is a standard shrinkage target in high-dimensional estimation, but in our application an interesting alternative target is to shrink the parameters towards the benchmark HAR parameters. To do this, we re-write the regularized regression models as:

$$RV_{t} = \beta_{0} + \widetilde{RV}_{t-1}(\beta_{HAR}^{d} + \gamma^{d}) + \frac{1}{4} \sum_{j=2}^{5} \widetilde{RV}_{t-j}(\beta_{HAR}^{w} + \gamma^{w}) + \frac{1}{16} \sum_{j=6}^{21} \widetilde{RV}_{t-j}(\beta_{HAR}^{m} + \gamma^{m}) + e_{t}$$
(S.1)

That is, we split the coefficients on high frequency returns into the HAR coefficients and a perturbation term to be estimated. Then, instead of regularizing the entire coefficient, we only penalize the perturbation terms, and shrink these towards zero. That is, the penalty term becomes:

$$\alpha(\lambda \|\boldsymbol{\gamma}\|_1 + (1-\lambda)\|\boldsymbol{\gamma}\|_2) \tag{S.2}$$

where $\gamma = [\gamma^d, \gamma^w, \gamma^m]$. The estimation procedure and hyperparameters grid are identical to those in the original regularized regression analysis. We compare the cubic HAR model with these alternative shrinkage estimators in Table S.1. (The OLS and HAR results are identical to those in Table 2 of the main paper, and are included here for ease of comparison.) With this alternative target the optimal degree of shrinkage is found to be large, and the Ridge, LASSO and elastic net models are all shrunk almost all the way to the benchmark HAR parameters. Given that, it is unsurprising that the cubic HAR continues to significantly out-perform these alternative regularized models, as shown in Table S.1, which have performance comparable to the benchmark HAR model.

Table S.1: Alternative regularized regression models: Shrink towards HAR

CubicHAR vs:	DM I Total	Losses Signif	DM Total	Wins Signif	DM t-stat Panel
Ridge- Alt	166	44	720	423	-17.3
LASSO-Alt	165	46	721	437	-4.2
$Elastic\ net ext{-}Alt$	170	46	716	412	-16.4
OLS	54	17	832	564	-6.4
HAR	155	41	731	470	-21.7

Note: This table reports individual and panel Diebold-Mariano tests comparing the cubic HAR model against competing models across 886 S&P 500 stocks. A positive panel DM t-statistic indicates that the competing model out-performs the cubic HAR model, while a negative t-statistic indicates the opposite. The models labeled "-Alt" shrink the estimated coefficients towards the HAR model coefficients, rather than towards zero as in Table 2 of the main paper.

S.4. Visualizing the optimal Cubic TOD and Cubic EW weights

This section presents the optimal daily weights for the cubic HAR, cubic TOD and cubic EW models, which differ in what shapes are estimated or imposed on the weekly and monthly weights. We observe the cubic TOD and cubic EW daily weights are almost identical to the cubic HAR weights for daily lags, which shows that restricting the shapes of weekly and monthly weights has almost zero affect on the estimated daily weights.

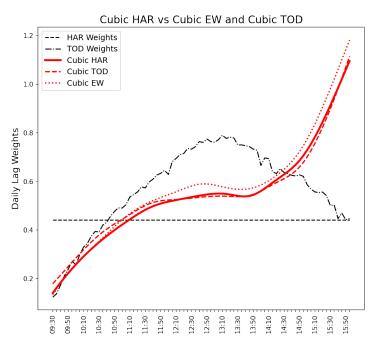


Figure S.2: Cubic TOD and Cubic EW weights

Note: This figure compares estimated cubic TOD and cubic EW weights against the cubic HAR model weights, averaged across the 886 S&P 500 stocks in our sample.

S.5. Forecast comparisons using mean-squared error loss

Table S.2: HAR vs other models performance on S&P500. MSE loss.

	DM Losses		\mathbf{DM}	Wins	DM t-stat
HAR vs:	Total	Signif	Total	Signif	Panel
Fully Flexible	421	169	465	106	0.8
$Flexible\ HAR$	726	408	160	11	3.3
$Cubic\ HAR$	714	406	172	7	1.9
$TOD\ HAR$	653	348	233	62	2.2
Ridge	462	186	424	120	-0.1
LASSO	203	46	683	338	-4.1
$Elastic\ net$	434	166	452	136	-2.0
OLS	64	7	822	507	-3.3

Note: This table reports individual and panel Diebold-Mariano tests comparing the baseline HAR model against competing models across 886 S&P 500 stocks, using MSE loss. A positive panel DM t-statistic indicates that the competing model out-performs the HAR model, while a negative t-statstic indicates the opposite. This table is related to Table 1 of the main paper, which uses QLIKE loss.

Table S.3: Cubic HAR vs other models performance on S&P 500. MSE loss.

	DM Losses		DM Wins		DM t-stat
Cubic HAR vs:	Total	Signif	Total	Signif	Panel
Fully Flexible	151	16	735	383	-0.7
Flexible HAR	411	55	475	90	0.0
$TOD\ HAR$	405	69	481	196	-1.0
HAR	172	7	714	406	-1.9
Ridge	171	12	715	349	-1.7
LASSO	66	2	820	542	-5.2
Elastic net	159	8	727	372	-3.4
OLS	23	3	863	664	-3.3

Note: This table reports individual and panel Diebold-Mariano tests comparing the cubic HAR model against competing models across 886 S&P 500 stocks, using MSE loss. A positive panel DM t-statistic indicates that the competing model out-performs the cubic HAR model, while a negative t-statistic indicates the opposite. This table is related to Table 2 of the main paper, which uses QLIKE loss.

S.6. Multi-days ahead forecasting comparisons: TOD HAR vs HAR

Table S.4: Multidays Ahead Volatility Forecasting: TOD HAR vs HAR

TOD vs HAR Horizon	DM Losses Total Signif		$egin{array}{ll} { m DM~Wins} \ { m Total} & { m Signif} \end{array}$		$egin{array}{l} ext{DM t-stat} \ ext{Panel} \end{array}$
1-Day	206	63	680	458	-18.6
2-Day	253	72	633	349	-12.5
β - Day	257	74	629	314	-10.1
4-Day	264	74	622	293	-13.6
5- Day	270	71	616	298	-11.5
20- Day	253	68	633	345	-6.8
60- Day	345	126	541	255	-0.2

Note: In this table, we report the individual and panel Diebold Mariano tests results of TOD HAR model against the HAR model in the S&P500 cross section for longer horizon forecasting. Note that negative test statistics favors the TOD HAR model.

S.7. Visualization of the hyperparameters for deep learning based models

In this section, we also report visualization of the cross sectional variation in the deep learning based models.

LR Epoch Batch 2048.0 100.0 52.1% 50.5% 0.01 99.1% 0.9% 0.1 10.2% 250.0 39.4% 47.9% 50.0 512.0 StepSize Gamma 0.25 0.1 54.6% 69.8% 9.5% 1.0 30.2% 35.9% 0.5 0.1

Figure S.3: Fully flexible model hyperparameters

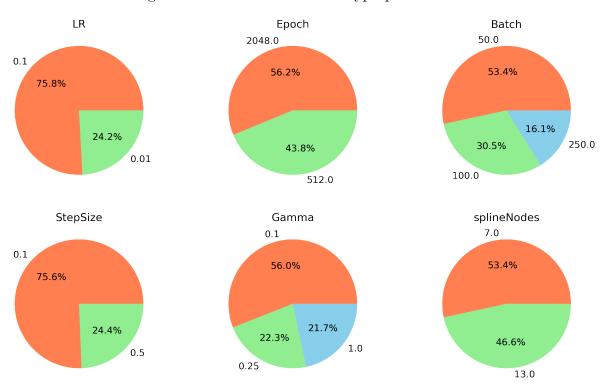
Note: This figure displays the cross-sectional variation of the optimal hyperparameters for the fully flexible model.

Figure S.4: Flexible HAR model hyperparameters



Note: This figure displays the cross-sectional variation of the optimal hyperparameters for the flexible HAR model.

Figure S.5: Cubic HAR model hyperparameters



Note: This figure displays the cross-sectional variation of the optimal hyperparameters for the cubic HAR model.

Figure S.6: Cubic-TOD model hyperparameters



Note: This figure displays the cross-sectional variation of the optimal hyperparameters for the cubic-TOD model.

Figure S.7: Cubic-EW model hyperparameters



Note: This figure displays the cross-sectional variation of the optimal hyperparameters for the cubic-EW model.